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# Decision Trees and Influence: an Inductive Proof of the OSSS Inequality

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**Abstract:** We give a simple proof of the OSSS inequality (O’Donnell, Saks, Schramm, Servedio, FOCS 2005). The inequality states that for any decision tree  $T$  calculating a Boolean function  $f : \{0, 1\}^n \rightarrow \{-1, 1\}$ , we have  $\text{Var}[f] \leq \sum_i \delta_i(T) \text{Inf}_i(f)$ , where  $\delta_i(T)$  is the probability that the input variable  $x_i$  is read by  $T$  and  $\text{Inf}_i(f)$  is the influence of the  $i$ th variable on  $f$ .

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## 1 Introduction

Let  $T$  be a decision tree computing a function  $f : \{0, 1\}^n \rightarrow \{-1, 1\}$ . We write  $\delta_i(T)$  for the probability that the variable  $x_i$  is queried by the decision tree on a uniform random input, and we write:

$$\Delta(T) \stackrel{\text{def}}{=} \sum_{i=1}^n \delta_i(T) = \mathbb{E}_{x \in \{0,1\}^n} [\# \text{ coordinates } T \text{ queries on } x].$$

$\Delta(T)$  can also be thought of as the average depth of the decision tree, or as a refinement of the notion of the size of the decision tree, since  $\Delta(T) \leq \log(\text{size}(T))$  [4].

The influence of a variable  $x_i$  on a Boolean function  $f$  is defined to be

$$\text{Inf}_i(f) = \Pr_{x \in \{0,1\}^n} [f(x) \neq f(x^{(i)})],$$

where  $x^{(i)}$  denotes  $x$  with its  $i$ -th bit flipped [1]. Recall that the variance of a function  $f$  is  $\text{Var}[f] = \mathbb{E}[(f - \mathbb{E}[f])^2]$ , and that the covariance of two functions  $f$  and  $g$  is  $\text{Cov}[f, g] = \mathbb{E}[(f - \mathbb{E}f)(g - \mathbb{E}g)]$ .

O’Donnell et al. [3] proved the following inequality:

**Theorem 1.1.** *Let  $f : \{0, 1\}^n \rightarrow \{-1, 1\}$  be a Boolean function, and let  $T$  be a decision tree computing  $f$ . Then*

$$\text{Var}[f] \leq \sum_{i=1}^n \delta_i(T) \text{Inf}_i(f).$$

This inequality can be viewed as a refinement of the Efron-Stein Inequality [2, 5] for the discrete cube (i. e.,  $\text{Var}[f] \leq \sum_{i=1}^n \text{Inf}_i(f)$ ) that takes into account the complexity of the function’s representation.

O’Donnell et al. [3] also proved the following generalization of [Theorem 1.1](#), which is what we reprove in the next section.

**Theorem 1.2.** *Let  $f, g : \{0, 1\}^n \rightarrow \{-1, 1\}$  be Boolean functions, and let  $T$  be a decision tree computing  $f$ . Then*

$$|\text{Cov}[f, g]| \leq \sum_{i=1}^n \delta_i(T) \text{Inf}_i(g).$$

## 2 Inductive Proof

The original proofs of both [Theorems 1.1](#) and [1.2](#) relied on some delicate probabilistic reasoning about the independence of certain hybrid inputs to the decision tree. We will prove [Theorem 1.2](#) using induction. To do so, we will consider the function’s behavior under the two cases when the root variable takes the value 0 and the value 1. First we will review a fact from probability theory.

For a function  $f : \{0, 1\}^n \rightarrow \mathbb{R}$ , let

$$c_i(x_1, \dots, x_n) = \mathbb{E}[f|(x_1, \dots, x_i)] - \mathbb{E}[f|(x_1, \dots, x_{i-1})]$$

for  $1 \leq i \leq n$  so that  $\mathbb{E}[f] + \sum_i c_i = f$ . The sequence  $\{c_i\}$  is a martingale difference sequence. Let  $g$  be another real-valued function, and let  $\{d_i\}$  be its martingale difference sequence. Then  $\text{Cov}[f, g] = \sum_{i=1}^n \mathbb{E}[c_i d_i]$ . We’ll prove this fact for the sake of completeness.

**Fact 2.1.** Let  $f, g : \{0, 1\}^n \rightarrow \mathbb{R}$  be real-valued functions, with martingale difference sequences

$$c_i = \mathbb{E}[f|(x_1, \dots, x_i)] - \mathbb{E}[f|(x_1, \dots, x_{i-1})] \quad \text{and} \quad d_i = \mathbb{E}[g|(x_1, \dots, x_i)] - \mathbb{E}[g|(x_1, \dots, x_{i-1})]$$

for  $1 \leq i \leq n$ . Then  $\text{Cov}[f, g] = \sum_{i=1}^n \mathbb{E}[c_i d_i]$ .

*Proof.* For  $j < k$ , we have

$$\mathbb{E}[c_j d_k] = \mathbb{E}[\mathbb{E}[c_j d_k|(x_1, \dots, x_{k-1})]] = \mathbb{E}[c_j \mathbb{E}[d_k|(x_1, \dots, x_{k-1})]] = \mathbb{E}[c_j \cdot 0] = 0.$$

Therefore  $\text{Cov}[f, g] = \mathbb{E}[(f - \mathbb{E}f)(g - \mathbb{E}g)] = \mathbb{E}[\sum_j c_j \sum_k d_k] = \sum_{i=1}^n \mathbb{E}[c_i d_i]$ . □

We now relate the last martingale difference sequence with the influence of the variable  $x_n$ .

**Lemma 2.2.** Let  $f, g : \{0, 1\}^n \rightarrow \{-1, 1\}$  be Boolean functions, and let

$$c_n = f - \mathbb{E}[f|(x_1, \dots, x_{n-1})] \quad \text{and} \quad d_n = g - \mathbb{E}[g|(x_1, \dots, x_{n-1})].$$

Then  $\mathbb{E}[c_n d_n] \leq \text{Inf}_n(f)$  and  $\mathbb{E}[c_n d_n] \leq \text{Inf}_n(g)$ .

*Proof.* Let  $f_0(x_1, \dots, x_n)$  denote  $f(x_1, \dots, x_{n-1}, 0)$ , and let  $f_1, g_0$ , and  $g_1$  be defined similarly. Then we have that

$$c_n = f - (f_0 + f_1)/2 \quad \text{and} \quad d_n = g - (g_0 + g_1)/2.$$

We can rewrite  $\mathbb{E}[c_n d_n]$  as

$$\mathbb{E}\left[\left(f - \frac{f_0 + f_1}{2}\right)\left(g - \frac{g_0 + g_1}{2}\right)\right].$$

If both  $f_0 \neq f_1$  and  $g_0 \neq g_1$ , the quantity inside the expectation is  $f \cdot g \in \{-1, 1\}$ . Otherwise, the quantity is 0.

The influence of  $x_n$  on  $f$  is just  $\text{Inf}_n(f) = \Pr[f_0(x) \neq f_1(x)]$ , and thus  $\text{Inf}_n(f)$  is an upper bound on  $\mathbb{E}[c_n d_n]$  (as is  $\text{Inf}_n(g)$ ). Note that this upper bound is an equality when we consider the special case of  $f = g$ , and we have that  $\mathbb{E}[c_n^2] = \text{Inf}_n(f)$ .  $\square$

We are now ready to prove [Theorem 1.2](#).

*Proof.* We'll prove the statement by induction on the number of variables. For the base case of one variable, recall that both  $\delta_i(T)$  and  $\text{Inf}_i(g)$  are always non-negative, and that the covariance of two functions with range  $\{-1, 1\}$  is a value in  $[-1, 1]$ . A Boolean function on only one variable is either constant, the single variable  $x_1$ , or its negation. If either  $f$  or  $g$  is constant, then  $\text{Cov}[f, g] = 0$ , and the inequality holds. If neither  $f$  nor  $g$  are constant, then  $\text{Inf}_1(g)$  and  $\delta_1(T)$  must be 1 and the inequality holds.

Now we'll consider  $f$  and  $g$  on  $n$  variables. We can assume that  $f$  and  $g$  are non-constant, or the inequality trivially holds as before. Thus,  $T$  must query at least one variable, and we will assume without loss of generality that the root of  $T$  queries  $x_n$ . Let  $T_0$  be the left subtree and let  $T_1$  be the right subtree. Then for  $i \neq n$ , we have  $\delta_i(T) = (1/2)\delta_i(T_0) + (1/2)\delta_i(T_1)$ . As in the proof of [Lemma 2.2](#), let  $f_0(x_1, \dots, x_n)$  denote  $f(x_1, \dots, x_{n-1}, 0)$ , and let  $f_1, g_0$ , and  $g_1$  be defined similarly. For  $i \neq n$ , we get the following expression:  $\text{Inf}_i(g) = (1/2)\text{Inf}_i(g_0) + (1/2)\text{Inf}_i(g_1)$ .

By [Fact 2.1](#), we can write  $\text{Cov}[f, g] = \sum_{i=1}^n \mathbb{E}[c_i d_i]$ . Let

$$c_{i,0} = \mathbb{E}[f_0|(x_1, \dots, x_i)] - \mathbb{E}[f_0|(x_1, \dots, x_{i-1})],$$

for  $1 \leq i \leq n-1$ , and define  $c_{i,1}, d_{i,0}$ , and  $d_{i,1}$  similarly. Then we have  $c_i = (c_{i,0} + c_{i,1})/2$ ,  $d_i = (d_{i,0} + d_{i,1})/2$ , and we can write the covariance as:

$$\text{Cov}[f, g] = \sum_{i=1}^n \mathbb{E}[c_i d_i] = \frac{1}{4} \sum_{i=1}^{n-1} \sum_{a,b \in \{0,1\}} \mathbb{E}[c_{i,a} d_{i,b}] + \mathbb{E}[c_n d_n] = \frac{1}{4} \sum_{a,b \in \{0,1\}} \text{Cov}[f_a, g_b] + \mathbb{E}[c_n d_n].$$

By the triangle inequality,  $|\text{Cov}[f, g]| \leq (1/4) \sum_{a,b \in \{0,1\}} |\text{Cov}[f_a, g_b]| + |\mathbb{E}[c_n d_n]|$ .

Since  $f_a$  and  $g_b$  are functions on  $n-1$  variables we can use the induction hypothesis, and we have:

$$|\text{Cov}[f, g]| \leq \frac{1}{4} \sum_{a, b \in \{0, 1\}} \sum_{i=1}^{n-1} \delta_i(T_a) \text{Inf}_i(g_b) + |\mathbb{E}[c_n d_n]| = \sum_{i=1}^{n-1} \delta_i(T) \text{Inf}_i(g) + |\mathbb{E}[c_n d_n]|.$$

As  $\text{Inf}_n(g) = \text{Inf}_n(-g)$ , we have that  $|\mathbb{E}[c_n d_n]| \leq \text{Inf}_n(g)$  by [Lemma 2.2](#), and  $\delta_n(T) = 1$  because  $x_n$  is the root of the tree. Thus, the inductive step holds.  $\square$

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